

3. $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Unif}(0, \theta)$ $\theta > 0$. Find $\hat{\theta}_{MLE}$.

$$L(\theta) = f(X_1, \dots, X_n; \theta) \stackrel{iid}{=} f(X_1; \theta) \cdots f(X_n; \theta)$$

$$f(X_i; \theta) = \begin{cases} \frac{1}{\theta} & \text{if } X_i \in [0, \theta] \\ 0 & \text{otherwise} \end{cases} = \frac{1}{\theta} \cdot I(X_i \leq \theta)$$

$$L(\theta) = \frac{1}{\theta} \cdot I(X_1 \leq \theta) \cdots \frac{1}{\theta} I(X_n \leq \theta)$$

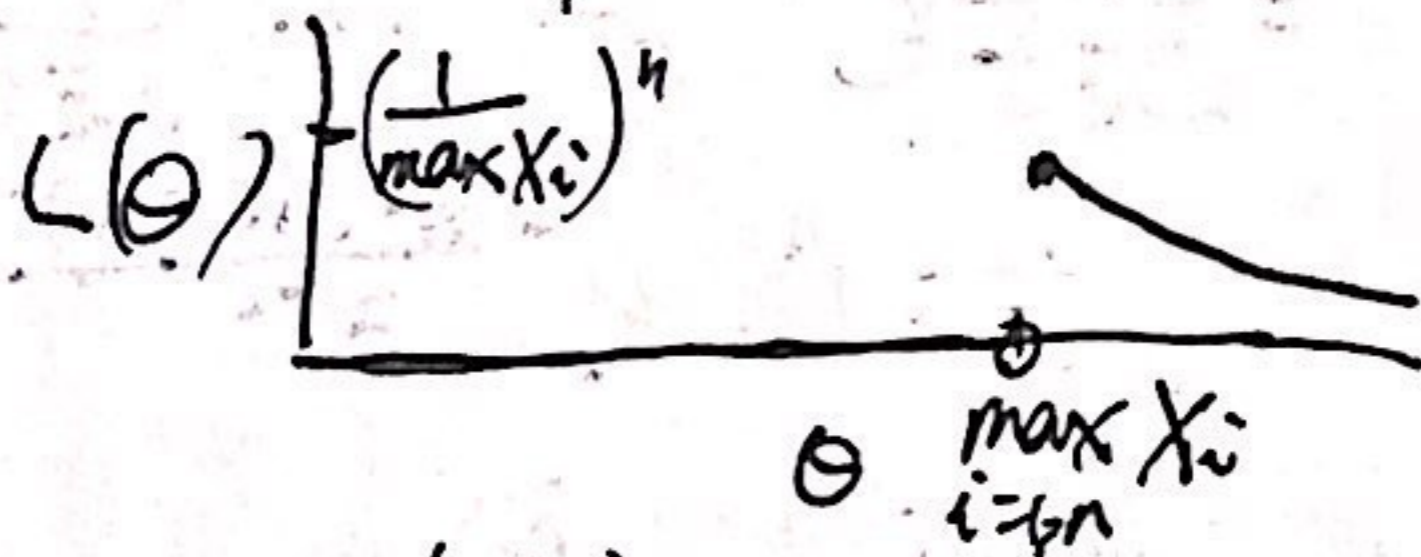
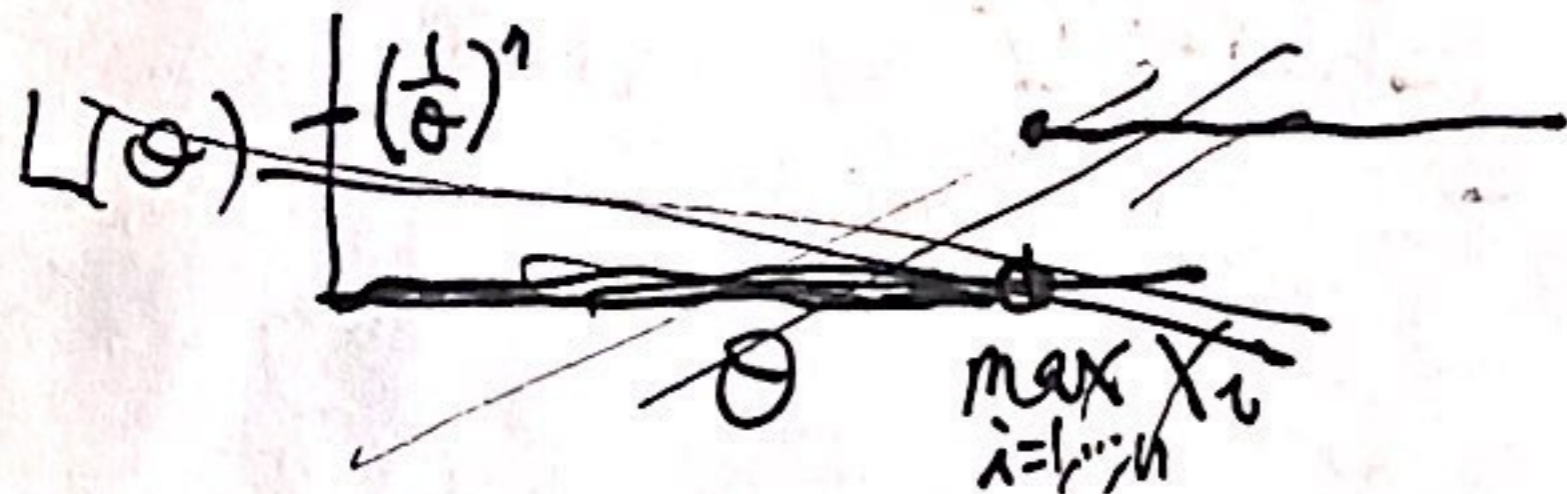
$$= \left(\frac{1}{\theta}\right)^n \cdot I(X_1 \leq \theta) \cdots I(X_n \leq \theta)$$

$$= \left(\frac{1}{\theta}\right)^n \cdot I(X_1 \leq \theta, \dots, X_n \leq \theta)$$

$$= \left(\frac{1}{\theta}\right)^n \cdot I\left(\max_{i=1, \dots, n} X_i \leq \theta\right) \Rightarrow \text{not differentiable with respect to } \theta$$

$$\hat{\theta}_{MOM} = 2\bar{X}_n$$

$$\hat{\theta}_{MLE} = \max_{i=1, \dots, n} X_i$$



$$\Rightarrow \arg \max_{\theta} L(\theta) = \max_{i=1, \dots, n} X_i$$

$$= \max\{X_1, \dots, X_n\}$$

Properties of $\hat{\theta}$ (estimator of θ)

① Bias: $\text{bias}(\hat{\theta}) = E[\hat{\theta}] - \theta$

② Variance: $\text{var}[\hat{\theta}] = E[(\hat{\theta} - E[\hat{\theta}])^2]$ (by def)

③ Mean-squared error (MSE): $\text{MSE}(\hat{\theta}) = E[(\hat{\theta} - \theta)^2]$ "how well does $\hat{\theta}$ concentrate on θ "

$$\text{MSE}(\hat{\theta}) = (\text{bias}(\hat{\theta}))^2 + \text{var}(\hat{\theta})$$

Two estimators $\hat{\theta}_1, \hat{\theta}_2$: if $\text{MSE}(\hat{\theta}_1) = \text{MSE}(\hat{\theta}_2)$, then

$$\text{bias}(\hat{\theta}_1) < \text{bias}(\hat{\theta}_2) \Leftrightarrow \text{var}(\hat{\theta}_1) > \text{var}(\hat{\theta}_2)$$

"bias-variance tradeoff"

Cramer-Rao Inequality (Theorem): the minimum variance of unbiased $\hat{\theta}$

Let $X_1, \dots, X_n \stackrel{iid}{\sim}$ pdf $f(\cdot; \theta)$. Let $\hat{\theta} = g(X_1, \dots, X_n)$ be an unbiased estimator of θ . Then under smoothness conditions of $f(\cdot; \theta)$,

$$\text{var}(\hat{\theta}) \geq \frac{1}{n I(\theta)} \quad \text{where } I(\theta) = \text{Fisher information}$$

Recall: $S_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2 \Rightarrow E[S_n^2] = \text{var}(X_i) = \sigma^2$ unbiased

$$\frac{n-1}{n} S_n^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2 \Rightarrow E\left[\frac{n-1}{n} S_n^2\right] = \frac{n-1}{n} E[S_n^2] = \frac{n-1}{n} \sigma^2$$

$$\Rightarrow \text{bias}\left(\frac{n-1}{n} S_n^2\right) = E\left[\frac{n-1}{n} S_n^2\right] - \sigma^2$$

So $\frac{n-1}{n} S_n^2$ is biased but asymptotically unbiased. $\frac{n-1}{n} \sigma^2 - \sigma^2 = -\frac{1}{n} \sigma^2 \xrightarrow{n \rightarrow \infty} 0$

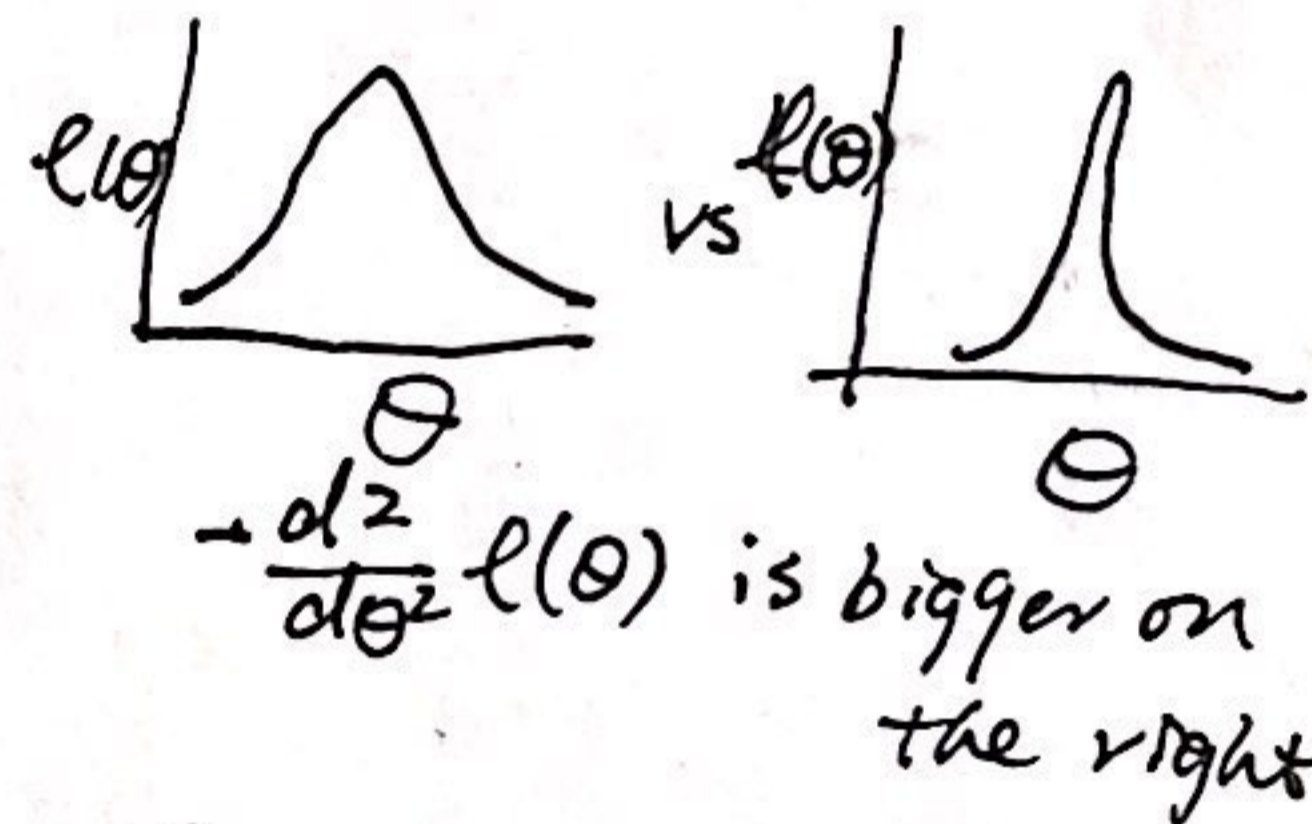
Fisher information: based on the log-likelihood of θ w/ 1 obs

$$\ell(\theta) = \log f(X; \theta) \quad f \text{ is the pdf or pmf of r.v. } X$$

Fisher information: $I(\theta) = E\left[\left(\frac{d}{d\theta} \ell(\theta)\right)^2\right]$ θ is a constant
↓
w.r.t. X

$$= E\left[\left(\frac{d}{d\theta} \log f(X; \theta)\right)^2\right]$$

Alternative def: $I(\theta) = -E\left[\frac{d^2}{d\theta^2} \ell(\theta)\right]$



Pf: Since $\int f(x; \theta) dx = 1$

$$\frac{d}{d\theta} \int f(x; \theta) dx = \frac{d}{d\theta} 1 = 0$$

Because of the smoothness conditions of $f(\cdot; \theta)$,

$$\frac{d}{d\theta} \int f(x; \theta) dx = \int \frac{d}{d\theta} f(x; \theta) dx = 0 \quad (1)$$

$$\text{Now, } \frac{d}{d\theta} \log f(x; \theta) = \frac{\frac{d}{d\theta} f(x; \theta)}{f(x; \theta)}$$

$$\text{so } \frac{d}{d\theta} f(x; \theta) = \left(\frac{d}{d\theta} \log f(x; \theta)\right) \cdot f(x; \theta)$$

Apply ~~Add~~ $\int \cdot dx$ to both

$$0 \stackrel{(1)}{=} \int \frac{d}{d\theta} f(x; \theta) dx = \int \left(\frac{d}{d\theta} \log f(x; \theta)\right) f(x; \theta) dx = E\left[\frac{d}{d\theta} \log f(X; \theta)\right]$$